

Robust ℓ_1 Approaches to Computing the Geometric Median and Principal and Independent Components

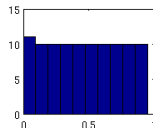
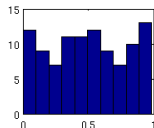
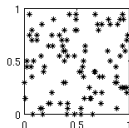
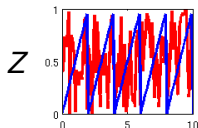
Stephen Keeling and Karl Kunisch
Institute for Mathematics and Scientific Computing
Karl Franzens University of Graz, Austria

Mathematical Image Processing Section
GAMM 2014, Erlangen
March 12, 2014



Graphical Demonstration of PCA/ICA

Sources Z , Measurements Y , sphered Y_s , separated X_s



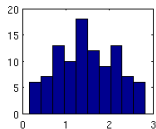
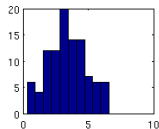
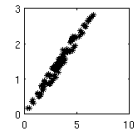
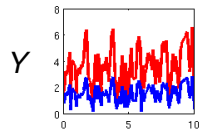
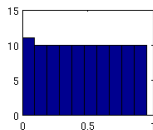
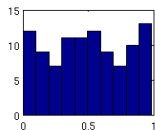
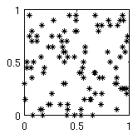
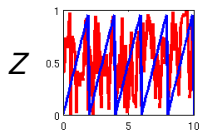
Time

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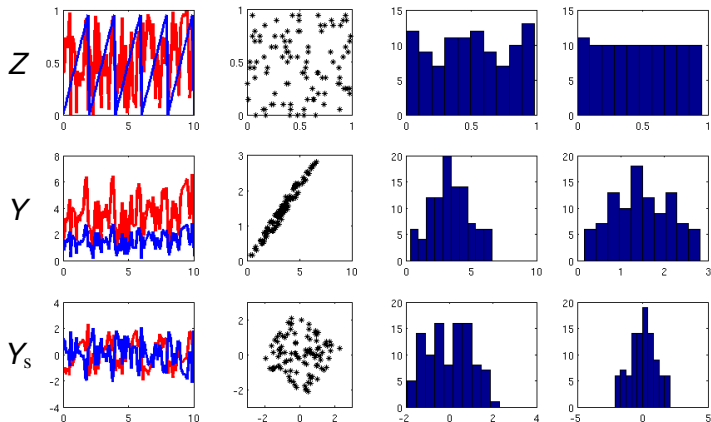
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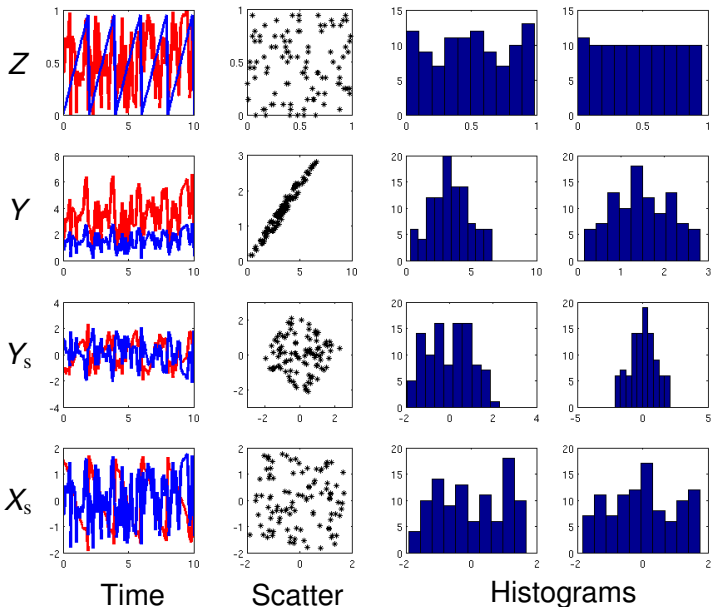
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Formulation of PCA/ICA

- Rows of Z are unknown samples of **sources** which are **independent and not Gauß distributed**.

$$Z = \begin{bmatrix} z_1(t_1) & z_1(t_2) & \cdots & z_1(t_n) \\ \vdots & \vdots & & \vdots \\ z_m(t_1) & z_m(t_2) & \cdots & z_m(t_n) \end{bmatrix}$$

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- ▶ Goal is to **undo** the trend toward **Gaussianity** to recover the sources

$$X = WY$$

with $W = U\Lambda^{-\frac{1}{2}}V^T \approx A^{-1}$ but unavoidable ambiguities.

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Summary of Steps:

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For example, **kurtosis**

$$\mathcal{K}(\mathbf{x}) = M_4(\mathbf{x}) - 3M_2^2(\mathbf{x})$$

satisfies $\mathcal{K}(\mathbf{n}) = 3\sigma^4 - 3\sigma^4 = 0$ for $\mathbf{n} \sim N(\mu, \sigma^2)$.

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So $J(\mathbf{u}) = [\mathcal{K}(Y_s^T \mathbf{u})]^2$ may be maximized with $\mathbf{u}_k^T \mathbf{u}_l = \delta_{kl}$.

Formulation of PCA/ICA

(PCA) Let the data be so decomposed,

$$Y_c = Y - \overline{Y}, \quad K = \frac{1}{n} Y_c Y_c^T, \quad KV = V\Lambda, \quad Y_s = \Lambda^{-\frac{1}{2}} V^T Y_c$$

Let $\Lambda = \text{diag}\{\lambda_1, \dots, \lambda_m\}$ with $\lambda_1 \geq \dots \geq \lambda_m$. With $P \in \mathbb{R}^{r \times m}$, $r < m$, $P_{i,j} = \delta_{i,j}$, the data Y are so projected to its ***r* strongest principal components**,

$$Y \approx Y_P = \overline{Y} + V\Lambda^{\frac{1}{2}} P^T P Y_s = \overline{Y} + \frac{1}{n} (P Y_s)^T (P Y_s)$$

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(ICA) Let the data be further so decomposed,

$$X_s = U Y_s$$

With $Q \in \mathbb{R}^{r \times m}$, $r < m$, $Q_{i,j} = \delta_{q_i,j}$, the data Y are so projected to the ***r* independent components** $\{q_1, \dots, q_r\}$,

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Generalization for higher dimensional data,

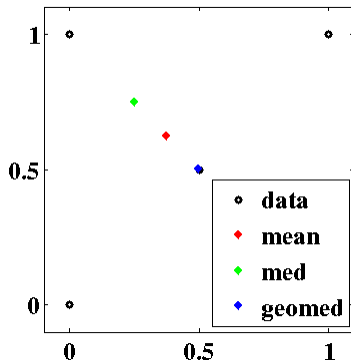
$$Y = \{\mathbf{y}_1, \dots, \mathbf{y}_n\} \in \mathbb{R}^{m \times n},$$

$$\text{geometric median}(Y) = \underset{\mu \in \mathbb{R}^m}{\operatorname{argmin}} M(\mu), \quad M(\mu) = \sum_{j=1}^n \|\mu - \mathbf{y}_j\|_{\ell_2}$$

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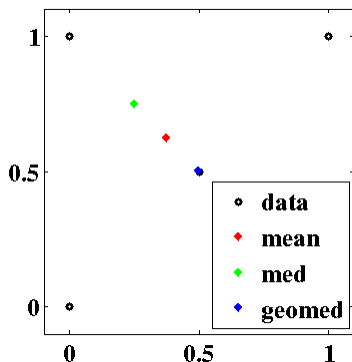


Data with Averages

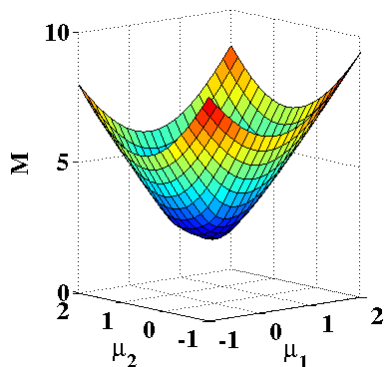
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Data with Averages



Geometric Median Landscape

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Sphered data:

$$Y_s = \Lambda^{-\frac{1}{2}} V^T Y_c, \quad Y_c = Y - \bar{Y} = \{\mathbf{y}_j^c\}$$

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$$\hat{\mathbf{v}}_1 = \underset{\|\hat{\mathbf{v}}\|_{\ell_2}=1}{\text{argmin}} \tilde{H}(\hat{\mathbf{v}})$$

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$$\text{where } \tilde{H}(\hat{\mathbf{v}}) = \sum_{j=1}^n \|(\hat{\mathbf{v}} \hat{\mathbf{v}}^T - I) \mathbf{y}_j^c\|_{\ell_2}^2$$

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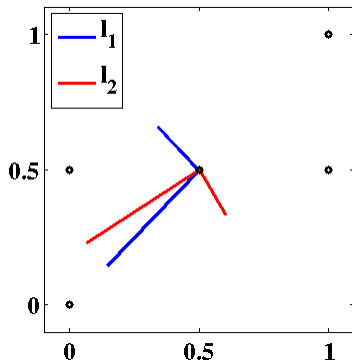
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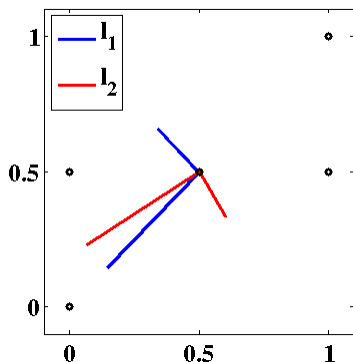


Data with Axes

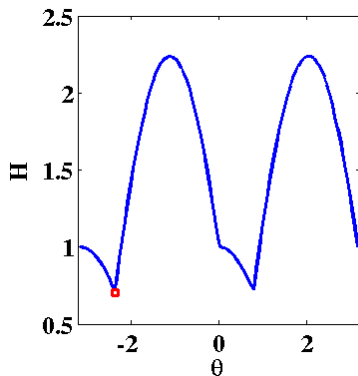
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Data with Axes



Rotation Landscape

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ℓ_2 formulation:

$$\hat{u}_1 = \operatorname{argmax}_{\|\hat{u}\|_{\ell_2}=1} K_\nu^2(\hat{u}^T Y_s) \quad \text{where} \quad K_\nu(\nu) = 0 \quad \text{for} \quad \nu \sim N(0, \cdot), \text{ e.g.,}$$

$$\begin{aligned} K_p(x) &= E[|x|^p] - \kappa_p E[x^2]^{p/2} \\ K_e(x) &= E[-\frac{1}{2}x^2] - 1/\sqrt{1 + E[x^2]} \end{aligned} \quad \kappa_p = \begin{cases} (p-1)!!, & p \text{ even} \\ \sqrt{\frac{2}{\pi}}(p-1)!!, & p \text{ odd} \end{cases}$$

and similarly in orthogonal complements for $\{\hat{u}_i\}_{i>1}$.

ℓ_1 formulation:

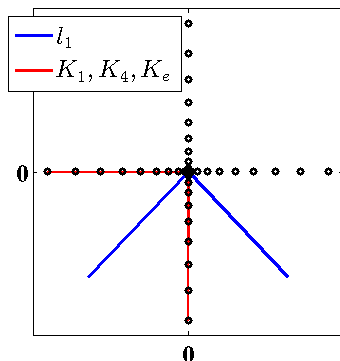
$$\hat{u}_1 = \operatorname{argmax}_{\|\hat{u}\|_{\ell_2}=1} G(\hat{u}) \quad \text{where} \quad G(\hat{u}) = \|\hat{u}^T Y_s\|_{\ell_1}$$

and similarly in orthogonal complements for $\{\hat{u}_i, \lambda_i\}_{i>1}$.

ℓ_2 and ℓ_1 Formulations of Independence

The data: $r = 20, k = 2, \sigma(t) = \text{sign}(t), i = -r, \dots, r,$

$$Y = \begin{bmatrix} -1 & \dots & \sigma(\frac{i}{r})|\frac{i}{r}|^k & \dots & 1 & 0 & \dots & \dots & 0 \\ 0 & \dots & \dots & \dots & 0 & -1 & \dots & \sigma(\frac{i}{r})|\frac{i}{r}|^k & \dots & 1 \end{bmatrix}$$

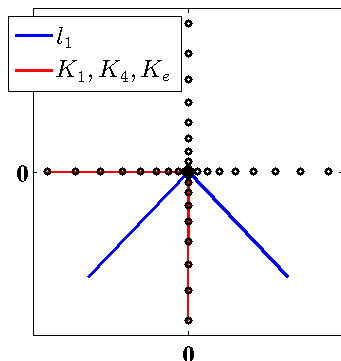


Data with Axes

ℓ_2 and ℓ_1 Formulations of Independence

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$$Y = \begin{bmatrix} -1 & \dots & \sigma(\frac{i}{r})|\frac{i}{r}|^k & \dots & 1 & 0 & \dots & \dots & 0 \\ 0 & \dots & \dots & \dots & 0 & -1 & \dots & \sigma(\frac{i}{r})|\frac{i}{r}|^k & \dots & 1 \end{bmatrix}$$



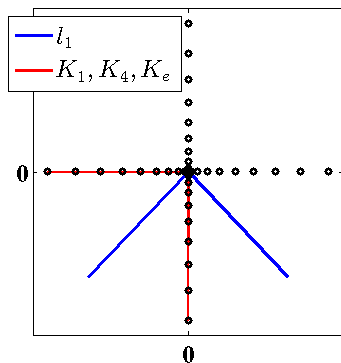
Data with Axes

$$P(x_1(\theta) = \alpha \text{ and } x_2(\theta) = \beta) \stackrel{?}{=} P(x_1(\theta) = \alpha) \cdot P(x_2(\theta) = \beta), \quad \forall \alpha, \beta \in \mathbb{R}$$

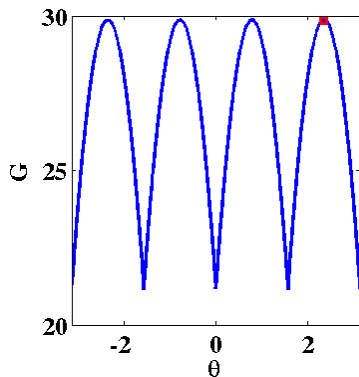
ℓ_2 and ℓ_1 Formulations of Independence

The data: $r = 20, k = 2, \sigma(t) = \text{sign}(t), i = -r, \dots, r,$

$$Y = \begin{bmatrix} -1 & \dots & \sigma(\frac{i}{r})|\frac{i}{r}|^k & \dots & 1 & 0 & \dots & \dots & 0 \\ 0 & \dots & \dots & \dots & 0 & -1 & \dots & \sigma(\frac{i}{r})|\frac{i}{r}|^k & \dots & 1 \end{bmatrix}$$



Data with Axes



Rotation Landscape

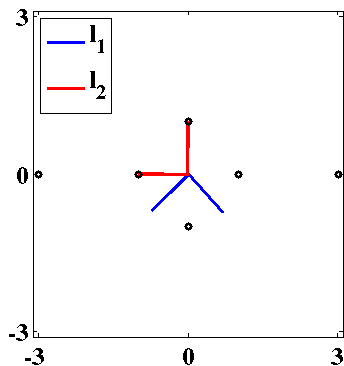
$$P(x_1(\theta) = \alpha \text{ and } x_2(\theta) = \beta) \stackrel{?}{=} P(x_1(\theta) = \alpha) \cdot P(x_2(\theta) = \beta), \quad \forall \alpha, \beta \in \mathbb{R}$$

ℓ_2 and ℓ_1 Formulations of Independence

The data:

$$Y_x = \begin{bmatrix} +1 & +1 & +1 & +1 & +1 & -1 & -1 & -1 & -1 & -1 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \end{bmatrix}$$

$$Y_y = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} Y_x, \quad Y_o = \begin{bmatrix} 3 & -3 \\ 0 & 0 \end{bmatrix}, \quad Y = \begin{bmatrix} Y_x & Y_y & Y_o \end{bmatrix}$$



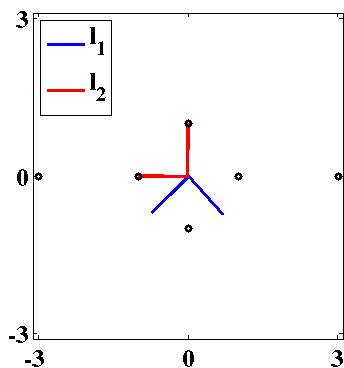
Data with Axes

ℓ_2 and ℓ_1 Formulations of Independence

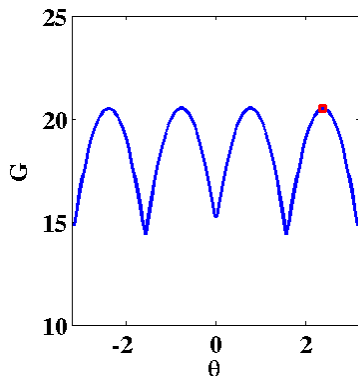
The data:

$$Y_x = \begin{bmatrix} +1 & +1 & +1 & +1 & +1 & -1 & -1 & -1 & -1 & -1 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \end{bmatrix}$$

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Data with Axes



Rotation Landscape

Numerical Optimization and Convergence

Geometric Median: $Y = \{\mathbf{y}_j\}$, $\mathbf{y}_j, \mathbf{d}_{j,l} \in \mathbb{R}^m$, $\boldsymbol{\mu}_l \in \mathbb{R}^m$, $\tau > 0$,

$$\mathbf{d}_{j,l+1} = \frac{\mathbf{d}_{j,l} + \tau(\boldsymbol{\mu}_l - \mathbf{y}_j)}{1 + \tau\|\boldsymbol{\mu}_l - \mathbf{y}_j\|_{\ell_2}}, \quad j = 1, \dots, n$$

$$\boldsymbol{\mu}_{l+1} = \sum_{j=1}^n \frac{(\mathbf{d}_{j,l} - \tau\mathbf{y}_j)}{1 + \tau\|\boldsymbol{\mu}_l - \mathbf{y}_j\|_{\ell_2}} / \sum_{j=1}^n \frac{-\tau}{1 + \tau\|\boldsymbol{\mu}_l - \mathbf{y}_j\|_{\ell_2}}.$$

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Theorem: Suppose the columns $\{\mathbf{y}_j\}_{j=1}^n = Y \in \mathbb{R}^{m \times n}$ are not colinear. Let $\{\mathbf{d}_1^*, \dots, \mathbf{d}_n^*, \boldsymbol{\mu}^*\}$ satisfy the necessary optimality condition for the geometric median merit function,

$$\boldsymbol{\mu}^* - \mathbf{y}_j = \|\boldsymbol{\mu}^* - \mathbf{y}_j\|_{\ell_2} \mathbf{d}_j^*, \quad \|\mathbf{d}_j^*\|_{\ell_2} \leq 1, \quad j = 1, \dots, n, \quad \sum_{j=1}^n \mathbf{d}_j^* = \mathbf{0}.$$

with $\boldsymbol{\mu}^* \notin \{\mathbf{y}_j\}_{j=1}^n$. Then $\{\mathbf{d}_1^*, \dots, \mathbf{d}_n^*, \boldsymbol{\mu}^*\}$ is a fixed point for the iteration which is locally asymptotically stable for τ large enough.

Numerical Optimization and Convergence

Given previous calculations,

$$Y_1 = Y_c, S_1 = \mathbb{R}^m$$

$$V_k = \{\hat{\mathbf{v}}_i\}_{i=1}^k, \quad Y_k = (I - V_{k-1} V_{k-1}^T) Y_c, \quad S_k = \mathcal{R}(V_{k-1})^\perp.$$

Sphering: $Y_k = \{\mathbf{y}_{j,k}\}$, $\mathbf{y}_{j,k}, \mathbf{d}_{j,l} \in S_k$, $\hat{\mathbf{v}}_l \in S_k$, $\tau, \rho > 0$

$$\mathbf{d}_{j,l+1} = \frac{(\hat{\mathbf{v}}_l \hat{\mathbf{v}}_l^T - I)(\tau \mathbf{y}_{j,k} - \mathbf{d}_{j,l})}{1 + \tau \|(\hat{\mathbf{v}}_l \hat{\mathbf{v}}_l^T - I) \mathbf{y}_{j,k}\|_{\ell_2}}, \quad j = 1, \dots, n$$

$$\hat{\mathbf{v}}_{l+1} = \frac{\mathbf{v}_{l+1}}{\|\mathbf{v}_{l+1}\|_{\ell_2}} \quad \text{with} \quad \mathbf{v}_{l+1} = \hat{\mathbf{v}}_l - \rho \sum_{j=1}^n \frac{(\hat{\mathbf{v}}_l^T \mathbf{y}_{j,k})(\tau \mathbf{y}_{j,k} - \mathbf{d}_{j,l})}{1 + \tau \|(\hat{\mathbf{v}}_l \hat{\mathbf{v}}_l^T - I) \mathbf{y}_{j,k}\|_{\ell_2}}.$$

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Theorem: Let $\{\mathbf{d}_1^*, \dots, \mathbf{d}_n^*, \hat{\mathbf{v}}^*\}$ satisfy for $1 \leq j \leq n$, $\|\mathbf{d}_j^*\|_{\ell_2} \leq 1$,

$$(\hat{\mathbf{v}}^* \hat{\mathbf{v}}^{*T} - I) \mathbf{y}_{j,k} = \|(\hat{\mathbf{v}}^* \hat{\mathbf{v}}^{*T} - I) \mathbf{y}_{j,k}\|_{\ell_2} \mathbf{d}_j^*, \quad \hat{\mathbf{v}}^{*T} \mathbf{d}_j^* = 0, \quad \sum_{j=1}^n (\hat{\mathbf{v}}^{*T} \mathbf{y}_{j,k}) \mathbf{d}_j^* = \mathbf{0}$$

with $\hat{\mathbf{v}}^* \in S_k$, $\|\hat{\mathbf{v}}^*\|_{\ell_2} = 1$, and suppose

$$\hat{\mathbf{v}}^{*T} \mathbf{y}_{j,k} \neq 0, \quad (\hat{\mathbf{v}}^* \hat{\mathbf{v}}^{*T} - I) \mathbf{y}_{j,k} \neq \mathbf{0}, \quad j = 1, \dots, n.$$

Then $\{\mathbf{d}_1^*, \dots, \mathbf{d}_n^*, \hat{\mathbf{v}}^*\}$ is a fixed point of the iteration which is locally asymptotically stable for τ large enough and ρ small enough.

Numerical Optimization and Convergence

Given previous calculations,

$$Y_1 = Y_s, T_1 = \mathbb{R}^m$$

$$U_l = \{\hat{\mathbf{u}}_i\}_{i=1}^l, \quad Y_l = (I - U_{l-1}^T U_{l-1}) Y_s, \quad T_l = \mathcal{R}(U_{l-1}^T)^\perp.$$

Independence: $\hat{\mathbf{u}}_k \in T_l, \tau > 0,$

$$\hat{\mathbf{u}}_{k+1} = \frac{\mathbf{u}_{k+1}}{\|\mathbf{u}_{k+1}\|_{\ell_2}} \quad \text{with} \quad \mathbf{u}_{k+1} = \hat{\mathbf{u}}_k + \tau [Y_l \sigma(Y_l^T \hat{\mathbf{u}}_k) - \hat{\mathbf{u}}_k \|\hat{\mathbf{u}}_k^T Y_l\|_{\ell_1}]$$

$$\text{where} \quad \begin{cases} \sigma(t) &= \text{sign}(t), & t \in \mathbb{R} \\ \sigma(\mathbf{v}) &= \{\sigma(v_j)\}_{j=1}^n, & \mathbf{v} = \{v_j\}_{j=1}^n \in \mathbb{R}^m. \end{cases}$$

Numerical Optimization and Convergence

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Theorem: Let $\hat{\mathbf{u}}^* \in T_l$ with $\|\hat{\mathbf{u}}^*\|_{\ell_2} = 1$ satisfy

$$Y_l \sigma(Y_l^T \hat{\mathbf{u}}^*) = \|\hat{\mathbf{u}}^*\|_{\ell_1} \hat{\mathbf{u}}^*$$

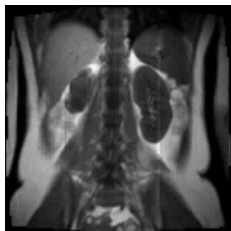
with $\mathcal{S} = \{j : \hat{\mathbf{e}}_j^T Y_l^T \hat{\mathbf{u}}^* = 0\} = \emptyset$. Then $\hat{\mathbf{u}}^*$ is a fixed point of the iteration which is locally asymptotically stable for τ small enough.

Application to DCE-MRI sequences

For each time $t = 1, \dots, T$, let the matrix of pixel values,

$$I(t) = \{I_{i,j}(t)\}_{1 \leq i,j \leq N}$$

be an image in a DCE-MRI sequence.



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Application to DCE-MRI sequences

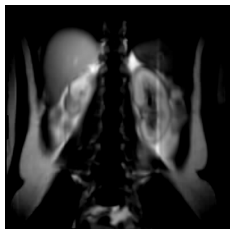
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With $n = T = 134$ and $m = N^2 = 400^2$ the images are represented as long vectors, and PCA/ICA is carried out with

$$Y = \{ \langle I_{1,1}(t), \dots, I_{N,1}(t), I_{1,2}(t), \dots, I_{N,2}(t), \dots, I_{1,N}(t), \dots, I_{N,N}(t) \rangle^T : t = 1, \dots, T \}$$



Application to DCE-MRI sequences

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To the **left is the first column of V** (displayed as an image),



Application to DCE-MRI sequences

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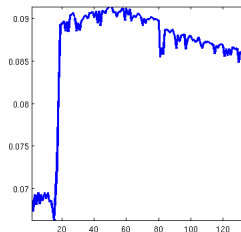
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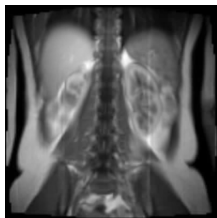
$$Y = \{ \langle I_{1,1}(t), \dots, I_{N,1}(t), I_{1,2}(t), \dots, I_{N,2}(t), \dots, I_{1,N}(t), \dots, I_{N,N}(t) \rangle^T : t = 1, \dots, T \}$$

To the **left is the first column of V** (displayed as an image), and to the **right is the first row of Y_s** .



Application to DCE-MRI sequences

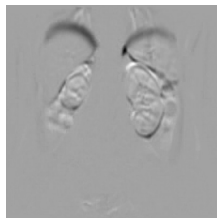
Basis: $\hat{\mathbf{v}}_i = V\hat{\mathbf{e}}_i$, $i = 1, 2$, $(\hat{\mathbf{e}}_i)_j = \delta_{ij}$,



$\bar{\mathbf{Y}}$



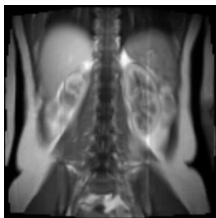
$\hat{\mathbf{v}}_1$



$\hat{\mathbf{v}}_2$

Application to DCE-MRI sequences

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\bar{Y}

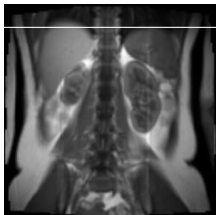
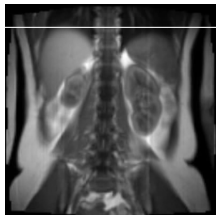


$\hat{\mathbf{v}}_1$



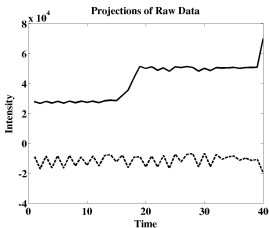
$\hat{\mathbf{v}}_2$

e.g., $\bar{Y} + V\Lambda^{\frac{1}{2}}\mathbf{e}_{ij}$, $i, j = \pm 1$, $\mathbf{e}_{ij} = (i, j, 0, \dots, 0)^T$

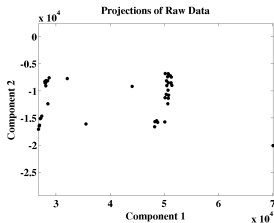


Application to DCE-MRI sequences

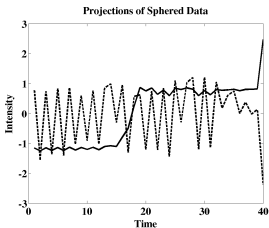
Decomposition with ℓ_2 PCA/ICA and outlier at $t = 40$:



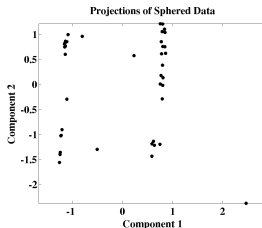
$$\hat{\mathbf{v}}_i^T \mathbf{Y} \text{ temporal}$$



$$\hat{\mathbf{v}}_i^T \mathbf{Y} \text{ phase}$$



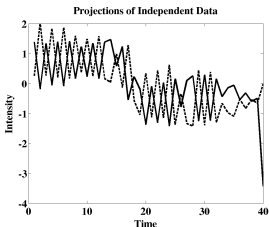
$$\hat{\mathbf{v}}_i^T \mathbf{Y}_s \text{ temporal}$$



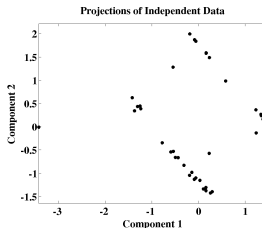
$$\hat{\mathbf{v}}_i^T \mathbf{Y}_s \text{ phase}$$

Application to DCE-MRI sequences

Decomposition with ℓ_2 PCA/ICA and outlier at $t = 40$:



$\hat{\mathbf{v}}_i^T \mathbf{X}_s$ temporal



$\hat{\mathbf{v}}_i^T \mathbf{X}_s$ phase



X_1

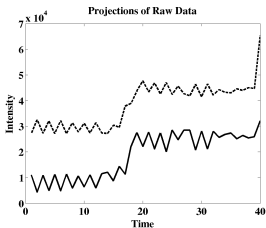


X_2

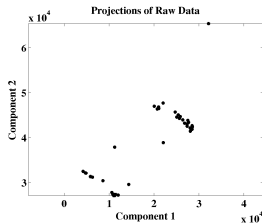
$$\mathbf{X}_i = \mathbf{V} \Lambda^{\frac{1}{2}} \mathbf{U}^T \mathbf{Q}_i \mathbf{X}_s, \mathbf{Q}_i = \text{diag}\{\hat{\mathbf{e}}_i\}, i = 1, 2$$

Application to DCE-MRI sequences

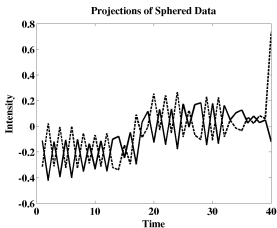
Decomposition with ℓ_1 PCA/ICA and outlier at $t = 40$:



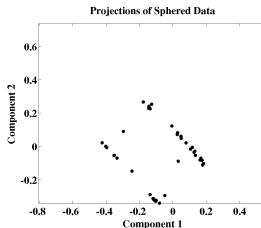
$\hat{\mathbf{v}}_i^T \mathbf{Y}$ temporal



$\hat{\mathbf{v}}_i^T \mathbf{Y}$ phase



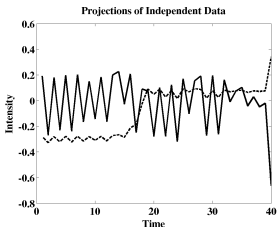
$\hat{\mathbf{v}}_i^T \mathbf{Y}_s$ temporal



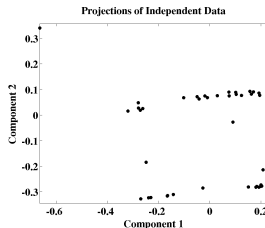
$\hat{\mathbf{v}}_i^T \mathbf{Y}_s$ phase

Application to DCE-MRI sequences

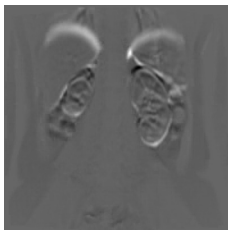
Decomposition with ℓ_1 PCA/ICA and outlier at $t = 40$:



$\hat{\mathbf{v}}_j^T \mathbf{X}_s$ temporal



$\hat{\mathbf{v}}_j^T \mathbf{X}_s$ phase



X_1

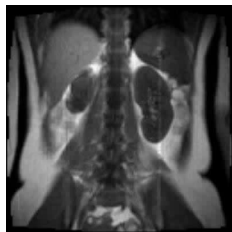


X_2

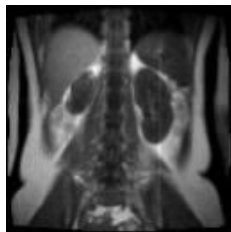
$$\mathbf{X}_i = \mathbf{V} \Lambda^{\frac{1}{2}} \mathbf{U}^T \mathbf{Q}_i \mathbf{X}_s, \quad \mathbf{Q}_i = \text{diag}\{\hat{\mathbf{e}}_i\}, \quad i = 1, 2$$

Application to DCE-MRI sequences

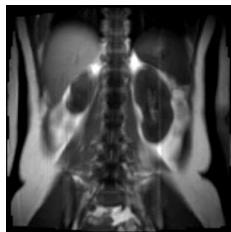
With motion component removed:



Original



Iterative Template



PCA/ICA

where Iterative Template sequence is given by iterating:

$$\begin{aligned}l_{\text{time}}(\mathbf{x}, \cdot) &= \underset{I}{\operatorname{argmin}} \int_0^T \left\{ |I - I_{\text{stat}}|^2 + \alpha |\partial_t I|^2 \right\} dt \quad (\text{init: } I_{\text{stat}} \leftarrow I_{\text{orig}}) \\ \mathbf{u}_{\text{space}}(\cdot, t) &= \underset{\mathbf{u}}{\operatorname{argmin}} \int_{\Omega} \left\{ |I_{\text{orig}} \circ (\text{Id} + \mathbf{u}) - I_{\text{time}}|^2 + \mu |\nabla \mathbf{u}^T + \nabla \mathbf{u}|^2 \right\} d\mathbf{x} \\ I_{\text{stat}} &\leftarrow I_{\text{orig}} \circ (\text{Id} + \mathbf{u}_{\text{space}})\end{aligned}$$

Application to DCE-MRI sequences

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Thank You!