## $\ell_{1}$ Approaches to PCA and ICA

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## Projekt: Signalverarbeitung

## Tutorium: Trennung von Datenquellen in unkorrelierte und unabhängige Komponenten

a.o.Univ.Prof. Mag.Dr. Stephen Keeling<br>http://math.uni-graz.at/keeling/

Literatur:
http://cis.legacy.ics.tkk.fi/aapo/papers/ IJCNN99_tutorialweb/
Dokumentation:
http://math.uni-graz.at/keeling/skripten/ Tutorium.pdf

Dank an Herrn Dipl.-Ing. Dr. Gernot Reishofer für seine Unterstützung für diese Arbeit!

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## Graphical Demonstration of PCA/ICA

Sources $Z$, Measurements $Y$, sphered $Y_{\mathrm{s}}$, separated $X_{\mathrm{c}}$



Time


Scatter


Histograms

## Formulation of PCA/ICA

- Rows of $Z$ are unknown samples of sources which are independent and not Gauß distributed.

$$
Z=\left[\begin{array}{cccc}
z_{1}\left(t_{1}\right) & z_{1}\left(t_{2}\right) & \cdots & z_{1}\left(t_{n}\right) \\
\vdots & \vdots & & \vdots \\
z_{m}\left(t_{1}\right) & z_{m}\left(t_{2}\right) & \cdots & z_{m}\left(t_{n}\right)
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- Rows of $Y$ are measured samples of unknown mixtures of the sources

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- Rows of $Y$ are measured samples of unknown mixtures of the sources

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no longer independent and now more Gauß distributed.

- Goal is to undo the trend toward Gaußianity to recover the sources

$$
X=W Y
$$

with $W=U \wedge^{-\frac{1}{2}} V^{\mathrm{T}} \approx A^{-1}$ but unavoidable ambiguity $x_{i}(t) \approx \pm z_{k}(t)$.

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- Centering:

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- Rotation:

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For example, Kurtosis

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\begin{gathered}
\mathcal{W}(\boldsymbol{x})=M_{4}(\boldsymbol{x})-3 M_{2}^{2}(\boldsymbol{x}) \\
\text { satisfies } \mathcal{W}(\boldsymbol{n})=3 \sigma^{4}-3 \sigma^{4}=0 \text { for } \boldsymbol{n} \sim N\left(\mu, \sigma^{2}\right)
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So $J(\boldsymbol{u})=-\mathcal{W}^{2}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)$ may be minimizied with $\boldsymbol{u}_{k}^{\mathrm{T}} \boldsymbol{u}_{I}=\delta_{k l}$.

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- Shift:

$$
X=X_{c}+U \Lambda^{-\frac{1}{2}} V^{\mathrm{T}} \bar{Y}=U \Lambda^{-\frac{1}{2}} V^{\mathrm{T}} Y=W Y
$$

## Formulation of PCA/ICA

(PCA) Let the data be so decomposed,

$$
Y_{\mathrm{c}}=Y-\bar{Y}, \quad K=\frac{1}{n} Y_{\mathrm{c}} Y_{\mathrm{c}}^{\mathrm{T}}, \quad K V=V \Lambda, \quad Y_{\mathrm{s}}=\Lambda^{-\frac{1}{2}} V^{\mathrm{T}} Y_{\mathrm{c}}
$$

Let $\Lambda=\operatorname{diag}\left\{\lambda_{1}, \ldots, \lambda_{m}\right\}$ with $\lambda_{1} \geq \cdots \geq \lambda_{m}$. With $P \in \mathbb{R}^{r \times m}$, $r<m, P_{i, j}=\delta_{i, j}$, the data $Y$ are so projected to its $r$ strongest principal components,

$$
Y \approx Y_{P}=\bar{Y}+V \wedge^{\frac{1}{2}} P^{\mathrm{T}} P Y_{\mathrm{s}}=\bar{Y}+\frac{1}{n}\left(P Y_{\mathrm{s}}\right)^{\mathrm{T}}\left(P Y_{\mathrm{s}}\right)
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$$

(ICA) Let the data be further so decomposed,

$$
X_{\mathrm{c}}=U Y_{\mathrm{s}}
$$

With $Q \in \mathbb{R}^{r \times m}, r<m, Q_{i, j}=\delta_{q_{i}, j}$, the data $Y$ are so projected to the $r$ independent components $\left\{q_{1}, \ldots, q_{r}\right\}$,

$$
Y \approx Y_{Q}=\bar{Y}+V \wedge^{\frac{1}{2}} U^{\mathrm{T}} Q^{\mathrm{T}} Q X_{\mathrm{c}}=\bar{Y}+\frac{1}{n}\left(Q X_{\mathrm{c}}\right)^{\mathrm{T}}\left(Q X_{\mathrm{c}}\right)
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\mu_{1}(\boldsymbol{x}) & =\arg \min _{\mu} \sum_{i=1}^{m}\left|\mu-x_{i}\right| \\
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Best generalization for higher dimensional data, $Y=\left\{\boldsymbol{y}_{1}, \ldots, \boldsymbol{y}_{m}\right\}^{\mathrm{T}} \in \mathbb{R}^{m \times n}$,

$$
\boldsymbol{\mu}_{1}(Y)=\arg \min _{\boldsymbol{\mu} \in \mathbb{R}^{n}} \sum_{i=1}^{m}\left\|\boldsymbol{\mu}-\boldsymbol{y}_{i}\right\|_{\ell_{2}}
$$

## Benefits of $\ell_{1}$ Formulations

Sphering. The $\ell_{2}$ approach is obtained by minimizing

$$
R_{k}(\boldsymbol{v})=\frac{\frac{1}{n}\left\langle Y_{k} Y_{k}^{\mathrm{T}} \boldsymbol{v}, \boldsymbol{v}\right\rangle}{\langle\boldsymbol{v}, \boldsymbol{v}\rangle}=\left[\frac{\left\|Y_{k}^{\mathrm{T}} \boldsymbol{v}\right\|_{\ell_{2}}}{\sqrt{n}\|\boldsymbol{v}\|_{\ell_{2}}}\right]^{2}
$$

where

$$
Y_{k}=\left(I-V_{k-1} V_{k-1}^{\mathrm{T}}\right) Y_{\mathrm{c}}, \quad k=2, \ldots, m-1, \quad Y_{1}=Y_{\mathrm{c}}
$$

and setting

$$
\boldsymbol{v}_{k}=\operatorname{argmin}_{\boldsymbol{v}} R_{k}(\boldsymbol{v}), \lambda_{k}=R_{k}\left(\boldsymbol{v}_{k}\right), V_{k}=\left\{\boldsymbol{v}_{1}, \ldots, \boldsymbol{v}_{k}\right\}, V=V_{m}
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Best generalization for $\ell_{1}$ is obtained by minimizing

$$
F_{k}(\boldsymbol{v})=\frac{\left\|Y_{k}^{\mathrm{T}} \boldsymbol{v}\right\|_{\ell_{1}}}{\sqrt{n}\|\boldsymbol{v}\|_{\ell_{2}}}
$$

where

$$
Y_{k}=\left(I-V_{k-1} V_{k-1}^{\mathrm{T}}\right) Y_{\mathrm{c}}, \quad k=2, \ldots, m-1, \quad Y_{1}=Y_{\mathrm{c}}
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## Benefits of $\ell_{1}$ Formulations

Outliers accumulated at $(0,1)$, then at $\left(0, \frac{1}{2}\right)$ and $\left(1, \frac{1}{2}\right)$,


Data and Their L1 and L2 Axes


Blue is for $\ell_{2}$, Red is for $\ell_{1}\left(\ell_{2}\right)$.

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Blue is for $\ell_{2}$, Red is for $\ell_{1}\left(\ell_{1}\right)$.

## Robust Objective Function for ICA

Test data $Y=\left\{\boldsymbol{y}_{1}, \boldsymbol{y}_{2}\right\}^{\mathrm{T}} \in \mathbb{R}^{2 \times n}$, each pair in $\{( \pm 1,0),(0, \pm 1)\}$ except for outliers

$$
\begin{gathered}
\left(\boldsymbol{y}_{1}\right)_{1}=\alpha, \quad\left(\boldsymbol{y}_{2}\right)_{1}=0 \\
\left(\boldsymbol{y}_{1}\right)_{2}=-\alpha, \quad\left(\boldsymbol{y}_{2}\right)_{2}=0
\end{gathered}
$$



Then $\bar{Y}=(0,0)$ and $V=I$.

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\begin{array}{cc}
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\end{array}
$$



Then $\bar{Y}=(0,0)$ and $V=I$.
The Kurtosis objective function $J(\boldsymbol{u})=-\mathcal{W}^{2}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)$ has the following landscape for the test data:

$\boldsymbol{u}=\{\cos (\theta), \sin (\theta)\}$ with $\theta=\frac{\pi}{4}$ is the robust solution.

This solution is obtained for $\alpha \approx 0$, but not for $\alpha$ moderately larger.

## Robust Objective Function for ICA

An alternative objection function is based on the $\ell_{1}$ moment,

$$
\mathcal{M}(\boldsymbol{x})=M_{1}(\boldsymbol{x})-\sqrt{M_{2}(\boldsymbol{x})} \sqrt{\frac{2}{\pi}}
$$

where $\mathcal{M}(\boldsymbol{n})=\sigma \sqrt{2 / \pi}-\sigma \sqrt{2 / \pi}=0$ for $\boldsymbol{n} \sim N\left(\mu, \sigma^{2}\right)$.

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The new objective function

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J(\boldsymbol{u})=-\mathcal{M}^{2}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)
$$

has the following landscape for the test data:


$$
\begin{aligned}
& \boldsymbol{u}=\{\cos (\theta), \sin (\theta)\} \text { with } \\
& \theta=\frac{\pi}{4} \text { is the robust solution. }
\end{aligned}
$$

This solution is obtained for a large range of $\alpha>0$.

## Minimizing the Robust Objective Function for ICA

The robust objective function

$$
J(\boldsymbol{u})=-\mathcal{M}^{2}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)=-\left[M_{1}\left(Y_{\mathrm{s}} \boldsymbol{u}\right)-\sqrt{2 / \pi}\right]^{2}
$$

$\left(M_{2}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)=1\right)$ is minimized under the condition $\boldsymbol{u}^{\mathrm{T}} \boldsymbol{u}=1$.

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$\left(M_{2}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)=1\right)$ is minimized under the condition $\boldsymbol{u}^{\mathrm{T}} \boldsymbol{u}=1$.
The solution is obtained from a stationary point of

$$
L(\boldsymbol{u}, \lambda)=-\left[M_{1}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)-\sqrt{2 / \pi}\right]^{2}+\lambda\left(\boldsymbol{u}^{\mathrm{T}} \boldsymbol{u}-1\right) / 2
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We have $D_{u} J(\boldsymbol{u})=-\phi(\boldsymbol{u}) G(\boldsymbol{u}) \boldsymbol{u}$ with

$$
\phi(\boldsymbol{u})=2\left[M_{1}\left(Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right)-\sqrt{2 / \pi}\right] \quad \text { and } \quad G(\boldsymbol{u})=\frac{1}{n} \sum_{i=1}^{n} \frac{Y_{\mathrm{s}} \boldsymbol{e}_{i} \boldsymbol{e}_{i}^{\mathrm{T}} Y_{\mathrm{s}}^{\mathrm{T}}}{\left|\boldsymbol{e}_{i}^{\mathrm{T}} Y_{\mathrm{s}}^{\mathrm{T}} \boldsymbol{u}\right|}
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## Minimizing the Robust Objective Function for ICA

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$$

A stationary point ( $\boldsymbol{u}^{\star}, \lambda^{\star}$ ) satisfies $-D_{\boldsymbol{u}} J\left(\boldsymbol{u}^{\star}\right)=\lambda^{\star} \boldsymbol{u}^{\star}$ or with $\lambda^{\star}=\mu^{\star}\left(\boldsymbol{u}^{\star}\right) \phi\left(\boldsymbol{u}^{\star}\right)$ the nonlinear eigenspace problem,

$$
\boldsymbol{G}\left(\boldsymbol{u}^{\star}\right) \boldsymbol{u}^{\star}=\mu^{\star}\left(\boldsymbol{u}^{\star}\right) \boldsymbol{u}^{\star}, \quad \boldsymbol{u}^{\star \mathrm{T}} \boldsymbol{u}^{\star}=1
$$

## Minimizing the Robust Objective Function for ICA

The nonlinear eigenspace problem is solved by a vector iteration.

Let $\boldsymbol{u}_{I} \approx \boldsymbol{u}^{\star}$ with $\left\|\boldsymbol{u}_{l}\right\|=1$ and an update $\boldsymbol{u}_{l+1}$ is determined by,

$$
\boldsymbol{u}=G\left(\boldsymbol{u}_{l}\right) \boldsymbol{u}_{l}, \quad \boldsymbol{u}_{l+1}=\boldsymbol{u} /\|\boldsymbol{u}\|, \quad l=1,2, \ldots
$$

After convergence

$$
\boldsymbol{u}^{\star}=\lim _{l \rightarrow \infty} \boldsymbol{u}_{I}
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is the first column of $U^{T}$.

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After convergence

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$$

is the first column of $U^{T}$.
The next column of $U^{T}$ is determined by a modified vector iteration.

For this, the projected data

$$
Y_{\mathrm{p}}=\left(I-\boldsymbol{u}^{\star} \boldsymbol{u}^{\star \mathrm{T}}\right) Y_{\mathrm{s}}
$$

have columns which are linearly independent from $\boldsymbol{u}^{\star}$.

## Minimizing the Robust Objective Function for ICA

 With the modified matrix,$$
\tilde{\boldsymbol{G}}(\boldsymbol{u})=\frac{1}{n} \sum_{i=1}^{n} \frac{Y_{\mathrm{p}} \mathbf{e}_{i} \boldsymbol{e}_{i}^{\mathrm{T}} Y_{\mathrm{p}}^{\mathrm{T}}}{\left|\boldsymbol{e}_{i}^{\mathrm{T}} Y_{\mathrm{p}}^{\mathrm{T}} \boldsymbol{u}\right|}
$$

the modified vector iteration is,

$$
\boldsymbol{u}=\left(I-\boldsymbol{u}^{\star} \boldsymbol{u}^{\star \mathrm{T}}\right) \tilde{G}\left(\boldsymbol{u}_{l}\right) \boldsymbol{u}_{l}, \quad \boldsymbol{u}_{l+1}=\boldsymbol{u} /\|\boldsymbol{u}\|, \quad I=1,2, \ldots
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## Minimizing the Robust Objective Function for ICA

 With the modified matrix,$$
\tilde{G}(\boldsymbol{u})=\frac{1}{n} \sum_{i=1}^{n} \frac{Y_{\mathrm{p}} \boldsymbol{e}_{i} \boldsymbol{e}_{i}^{\mathrm{T}} Y_{\mathrm{p}}^{\mathrm{T}}}{\left|\boldsymbol{e}_{i}^{\mathrm{T}} Y_{\mathrm{p}}^{\mathrm{T}} \boldsymbol{u}\right|}
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the modified vector iteration is,

$$
\boldsymbol{u}=\left(I-\boldsymbol{u}^{\star} \boldsymbol{u}^{\star \mathrm{T}}\right) \tilde{G}\left(\boldsymbol{u}_{l}\right) \boldsymbol{u}_{l}, \quad \boldsymbol{u}_{I+1}=\boldsymbol{u} /\|\boldsymbol{u}\|, \quad I=1,2, \ldots
$$

The remaining columns of $U^{\mathrm{T}}$ are determined similarly, where $\boldsymbol{u}^{\star}$ above is replaced with the matrix $\left[\boldsymbol{u}_{1}^{\star}, \ldots, \boldsymbol{u}_{k}^{\star}\right]$, when $k$ columns $\left\{\boldsymbol{u}_{1}^{\star}, \ldots, \boldsymbol{u}_{k}^{\star}\right\}$ of $U^{\mathrm{T}}$ have already been calculated.

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Observation: The vector iterations converge very robustly to the global constrained minimum.

Claim: At least convergence to a local constrained minimum can be proved with adequate step size control.

## Application to DCE-MRI sequences

For each time $t=1, \ldots, T$, the matrix of pixel values,

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B(t)=\left\{B_{i, j}(t)\right\}_{1 \leq i, j \leq N}
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With $m=T=134$ and $n=N^{2}=400^{2}$ the images are represented as long vectors:

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\boldsymbol{y}_{t}^{\mathrm{T}}=\left\{B_{1,1}(t), \ldots, B_{N, 1}(t), B_{1,2}(t), \ldots, B_{N, 2}(t), \ldots, B_{1, N}(t), \ldots, B_{N, N}(t)\right\}
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and PCA/ICA is carried out with $Y^{\mathrm{T}}=\left\{\boldsymbol{y}_{1}, \ldots, \boldsymbol{y}_{m}\right\}$.
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( $Y \leftarrow Y^{\mathrm{T}}$ is equivalent but expensive.)
To the left is the first row of $Y_{\mathrm{s}}$ (displayed as image),



To the right is the first column of $V$.

## Application to DCE-MRI sequences

Top 6 independent components:


Eliminate motion, keep anatomy and contrast agent: [Video]

## Application to DCE-MRI sequences

Virtual Gating, through segmentation of correlations:


Three groups [Video], stabilized further by PCA/ICA [Video].

## Formulation in Function Space

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Consequence: The function space setting resembles the finite dimensional setting but with infinite matrices operating between bases in separable spaces.

